# Lifecycle Residential Mobility and the Intergenerational Persistence of Economic Status 

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## Segment I: <br> Childhood Exposure Effects: <br> Causality vs Selection and Sorting

## Introduction

- Chetty et al. (2018) analyze data on 7 million families who moved across commuting zones (CZ) in the US

■ They conclude that neighborhoods (nbhd) in which children grow up shape their various adulthood outcomes

- Outcomes of children whose families move to a better nbhd -measured by outcomes of children already living thereimprove in proportion to the amount of time they spend growing up in that nbhd, at a rate of $4 \%$ per year of exposure
- The main identification in Chetty et al. (2018) is using variation in the age of children when families move
- They interpret their results as causal effects of neighborhoods
- Robustness checks using variations among siblings, displacement shocks etc.


## Introduction- Cont'd

■ We re-examine Chetty et al. (2018) in a Danish context

- First, compare our results to US results in Chetty et al. (2018)
- Then, investigate the mechanisms behind our results
- Can one interpret the results as causal effects of neighborhoods or "power of place"?
- The role of selection and sorting
- A core identifying assumption in Chetty et al. (2018): selection effects do not vary with the child's age at move
- This means that children potential outcomes are orthogonal to their age when their parents move across neighborhoods


## Preview of Results

- We find similar estimates to those of Chetty et al. (2018) but with a lower precision as we utilize Danish population data

■ We provide evidence for violation of main identifying assumption in previous works (constant-in-age selection effect)

- Three sets of results pointing to violation of constant-in-age selection effect:

■ Self-selection into "permanent residency" status and self-selection into timing of moves (wrt age of children)

- When moving, people sort into nbhds and the age of child when parents move is not orthogonal to the extent to which there is a positive sorting between parents and neighborhoods:
- Higher quality moves in early childhood
- Higher correlation of later moves with income/family shocks


# Section 1: Literature Review 

## Challenges and Questions

- What do we learn from previous works about the role of nbhd?
- nbhd char.: schools, crime, housing stock, air quality, etc
- Measurement errors
- nbhd quality defined solely based on children of PRs' inc. rank
- static nbhd quality
- measure of resources (family unit, inc type, transitory shocks)
- definition of PRs and movers
- missing early years of childhood (before age 9)
- External validity:
- dynamic results may not be extrapolated to early childhood
- not clear implication for non-movers: identification
- Methodology:
- rank-rank analysis
- welfare implications
- lack of a life-cycle approach

■ Identifying assumptions

- complementarity between early- and late-childhood investments
- constant-in-age selection

Chetty et al. (2018):
THE IMPACTS OF NEIGHBORHOODS ON INTERGENERATIONAL MOBILITY I: CHILDHOOD EXPOSURE EFFECTS

## Data

■ Data source: Federal income tax records

- Data span: 1996-2012

■ Sample: Children who were born between 1980-1988 (covering ages 8-24 of the last cohort \& and 16-32 of first)

- permanent residents (stayers/PR): subset of parents who reside in a single CZ c in 1996-2012. (not robust)
- movers: individuals in the main sample who are not PR

■ Income type: Adjusted gross inc. (1040 tax return) + tax-exempt interest inc. and the nontaxable SSDI benefits

- averaged over 1996-2000 to get parent inc; age 24 for child

■ Unit of Analysis: Family income
■ Estimation Sample: Only PR and those who moved across nbhds exactly once during 1996-2012 (ad-hoc)

TABLE I
Summary Statistics for CZ Permanent Residents and Movers

|  | Mean | Std. dev. | Median | Num. of obs. |
| :--- | :---: | :---: | :---: | :---: |
| Variable | (1) | (2) | (3) | (4) |

Panel A: Permanent residents: Families who do not move across CZs

| Parent family income | 89,909 | 357,194 | 61,300 | $19,499,662$ |
| :--- | :---: | :---: | :---: | :---: |
| Child family income at 24 | 24,731 | 140,200 | 19,600 | $19,499,662$ |
| Child family income at 26 | 33,723 | 161,423 | 26,100 | $14,894,662$ |
| Child family income at 30 | 48,912 | 138,512 | 35,600 | $6,081,738$ |
| Child individual income at 24 | 20,331 | $\mathbf{1 3 9 , 6 9 7}$ | 17,200 | $19,499,662$ |
| Child married at 26 | 0.25 | 0.43 | 0.00 | $12,997,702$ |
| Child married at 30 | 0.39 | 0.49 | 0.00 | $6,081,738$ |
| Child attends college between 18-23 | 0.70 | 0.46 | 1.00 | $17,602,702$ |
| Child has teen birth (females only) | 0.11 | 0.32 | 0.00 | $9,670,225$ |
| Child working at age 16 | 0.41 | 0.49 | 0.00 | $13,417,924$ |

Panel B: Families who move 1-3 times across CZs

| Parent family income | 90,468 | 376,413 | 53,500 | $4,374,418$ |
| :--- | :---: | :---: | :---: | :---: |
| Child family income at 24 | 23,489 | 57,852 | 18,100 | $4,374,418$ |
| Child family income at 26 | 31,658 | 99,394 | 23,800 | $3,276,406$ |
| Child family income at 30 | 46,368 | 107,380 | 32,500 | $1,305,997$ |
| Child individual income at 24 | 19,091 | 51,689 | 15,600 | $4,374,418$ |
| Child married at 26 | 0.25 | 0.43 | 0.00 | $2,867,598$ |
| Child married at 30 | 0.38 | 0.49 | 0.00 | $1,305,997$ |
| Child attends college between 18-23 | 0.66 | 0.47 | 1.00 | $3,965,610$ |
| Child has teen birth (females only) | 0.13 | 0.33 | 0.00 | $2,169,207$ |
| Child working at age 16 | 0.40 | 0.49 | 0.00 | $3,068,421$ |

Panel C: Primary analysis sample: families who move exactly once across CZs

| Parent family income | 97,064 | 369,971 | 58,700 | $1,553,021$ |
| :--- | :---: | :---: | :---: | :---: |
| Child family income at 24 | 23,867 | 56,564 | 18,600 | $1,553,021$ |
| Child family income at 26 | 32,419 | 108,431 | 24,500 | $1,160,278$ |
| Child family income at 30 | 47,882 | 117,450 | 33,600 | 460,457 |
| Child individual income at 24 | 19,462 | 48,452 | 16,000 | $1,553,021$ |
| Child married at 26 | 0.25 | 0.43 | 0.00 | $1,016,264$ |
| Child married at 30 | 0.38 | 0.49 | 0.00 | 460,457 |
| Child attends college between 18-23 | 0.69 | 0.46 | 1.00 | $1,409,007$ |
| Child has teen birth (females only) | 0.11 | 0.32 | 0.00 | 769,717 |
| Child working at age 16 | 0.39 | 0.49 | 0.00 | $1,092,564$ |

## Geographical Variation in Outcomes of PR

- Given birth cohort s and CZ c , let p be the parents' percentile in the national income distribution

■ Let $y_{i}$ denote the child's national income rank in adulthood

## Geographical Variation in Outcomes of PR- Cont'd



Figure I
Mean Child Income Rank versus Parent Income Rank for Children Raised in Chicago

## Geographical Variation in Outcomes of PR

$$
y_{i}=\alpha_{c s}+\psi_{c s} p_{i}+\epsilon_{i}
$$

then, estimate $y_{p c s}$, the mean rank of children with parents at percentile $p$ of the income distribution in CZ c in birth cohort s , using the fitted values:

$$
\bar{y}_{p c s}=\hat{\alpha}_{c s}+\hat{\psi}_{c s} p
$$

For example, $\bar{y}_{25, c, 1980}=40.1$ for children growing up at the 25 th percentile of the national income distribution and $\bar{y}_{75, c, 1980}=59.3$ for children growing up at the 75th percentile.

## Mean Inc. Ranks for Children with Parents at $25^{\text {th }}$ Pctile

(A) For Children with Parents at the $25^{\text {th }}$ Percentile


## Exposure Effects

Exposure effect at age $m$ : the impact of spending year $m$ of one's childhood in an area where PR's outcomes are 1 pp higher

Thought experiment: randomly assign children to new nbhd $d$ starting at age $m$ for the rest of childhood. The best linear predictor of children's outcomes $y_{i}$ in the experimental sample, based on the PR's outcomes in CZ $d\left(\bar{y}_{p d s}\right)$ :

$$
\begin{equation*}
y_{i}=\alpha_{m}+\beta_{m} \bar{y}_{p d s}+\theta_{i} \tag{3}
\end{equation*}
$$

Random assignment: $\theta \perp \bar{y}_{p d s}$
Exposure effect at $m: \gamma_{m}=\beta_{m}-\beta_{m+1}$, the effect on $y_{i}$ of spending the year from age m to age $\mathrm{m}+1$ in the destination Observational data: $b_{m}=\beta_{m}+\delta_{m}$
Bias $=\delta_{m}=\frac{\operatorname{cov}\left(\theta_{i}, \bar{y}_{p d s}\right)}{\operatorname{var}\left(\bar{y}_{p d s}\right)}$ : parent inputs \& unobserved det. of children's outcomes covary with PR's outcomes

## Exposure Effects- Constant-in-Age Selection Assumption

Bias $=\delta_{m}=\frac{\operatorname{cov}\left(\theta_{i}, \bar{y}_{p d s}\right)}{\operatorname{var}\left(\bar{y}_{p d s}\right)}$
ASSUMPTION 1 (A.1): Selection effects do not vary with the child's age at move: $\delta_{m}=\delta$ for all m.

Under A.1, we obtain consistent estimates of exposure effects:

$$
\gamma_{m}=\left(\beta_{m}+\delta_{m}\right)-\left(\beta_{m+1}+\delta_{m+1}\right)=b_{m}-b_{m+1}
$$

Even in observational data because the selection effects $\delta$ cancel out when estimating the exposure effect.

Rules out differential preferences among parents by age of child for local amenities, such as school quality, that are not fully captured in adult income percentile rank $\bar{y}_{p d s}$

## Exposure Effects- Estimation

To begin, consider the set of children whose families moved when they were exactly $m$ years old.

We analyze how these children's incomes in adulthood are related to those of PR in their destination CZ using the following linear regression:

$$
\begin{equation*}
y_{i}=\alpha_{q o s}+b_{m} \Delta_{o d p s}+\epsilon_{1 i} \tag{4}
\end{equation*}
$$

where $y_{i}$ denotes the child's income rank at age 24, $\alpha_{\text {qos }}$ is a fixed effect for the origin CZ o by parent income decile $q$ by birth cohort $s$ and $\Delta_{o d p s}=\bar{y}_{p d s}-\bar{y}_{p o s}$ is the difference in predicted income rank (at age 24) of permanent residents in the destination versus origin for the relevant parent income rank p and birth cohort $s$.

## Movers' Outcomes versus Predicted Outcomes Based on PR in Destination- Movers at Age 13



## Childhood Exposure Effects on Inc. Ranks in Adulthood

(5)

$$
y_{i}=\alpha_{q o s m}+\sum_{m=9}^{30} b_{m} I\left(m_{i}=m\right) \Delta_{o d p s}+\sum_{s=1980}^{1987} \kappa_{s} I\left(s_{i}=s\right) \Delta_{o d p s}+\varepsilon_{2 i},
$$

$\Delta_{\text {qosm }}$ : (origin $\times$ parent income decile $\times$ birth cohort $\times$ age) FE $\hat{b}_{m}$ : the average effect on age- 24 income rank $y_{i}$, conditional on moving from $o$ to $d$ at age $m$, of a 1 percentile increase in $\Delta_{o d p s}$

## Childhood Exposure Effects on Inc. Ranks in Adulthood

(5)

$$
y_{i}=\alpha_{q o s m}+\sum_{m=9}^{30} b_{m} I\left(m_{i}=m\right) \Delta_{o d p s}+\sum_{s=1980}^{1987} \kappa_{s} I\left(s_{i}=s\right) \Delta_{o d p s}+\varepsilon_{2 i}
$$

$\Delta_{\text {qosm }}$ : (origin $\times$ parent income decile $\times$ birth cohort $\times$ age) FE $\hat{b}_{m}$ : the average effect on age- 24 income rank $y_{i}$, conditional on moving from $o$ to $d$ at age $m$, of a 1 percentile increase in $\Delta_{o d p s}$ Alternative: parametric model estimating cohort- and age-specific slopes instead of FE

$$
y_{i}=\sum_{s=1980}^{1988} I\left(s_{i}=s\right)\left(\alpha_{s}^{1}+\alpha_{s}^{2} \bar{y}_{p o s}\right)+\sum_{m=9}^{30} I\left(m_{i}=m\right)\left(\zeta_{m}^{1}+\zeta_{m}^{2} p_{i}\right)
$$

$$
\begin{equation*}
+\sum_{m=9}^{30} b_{m} I\left(m_{i}=m\right) \Delta_{o d p s}+\sum_{s=1980}^{1987} \kappa_{s}^{d} I\left(s_{i}=s\right) \Delta_{o d p s}+\varepsilon_{3 i} \tag{6}
\end{equation*}
$$

## Results: $\hat{b}_{m}$ as Function of Age $m$



Figure IV
Childhood Exposure Effects on Income Ranks in Adulthood

## Results: $\hat{b}_{m}$ as Function of Age m- Parametric Estimates

(B) Parametric Estimates


Figure IV
Childhood Exposure Effects on Income Ranks in Adulthood

## Childhood Exposure Effect Estimates- Specification

$$
\begin{aligned}
y_{i}= & \sum_{s=1980}^{1988} I\left(s_{i}=s\right)\left(\alpha_{s}^{1}+\alpha_{s}^{2} \bar{y}_{p o s}\right)+\sum_{m=9}^{30} I\left(m_{i}=m\right)\left(\zeta_{m}^{1}+\zeta_{m}^{2} p_{i}\right) \\
& +\sum_{s=1980}^{1987} \kappa_{s}^{d} I\left(s_{i}=s\right) \Delta_{o d p s}+I\left(m_{i} \leqslant 23\right)\left(b_{0}+\left(23-m_{i}\right) \gamma\right) \Delta_{o d p s} \\
(7) \quad & +I\left(m_{i}>23\right)\left(\delta+\left(23-m_{i}\right) \delta^{\prime}\right) \Delta_{o d p s}+\varepsilon_{3 i} .
\end{aligned}
$$

## Childhood Exposure Effect Estimates- Results

TABLE II
Childhood Exposure Effect Estimates

| Specification: | Dependent variable: Child's income rank at age 24 |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  | With | amily fixed | ffects |
|  | Pooled <br> (1) | $\text { Age } \leqslant 23$ <br> (2) | $\text { Age }<18$ <br> (3) | No cohort controls <br> (4) | Individual income <br> (5) | Child CZ FE <br> (6) | Baseline <br> (7) | No cohort controls (8) | Timevarying controls (9) |
| Exposure effect ( $\gamma$ ) | $\begin{gathered} 0.040 \\ (0.002) \end{gathered}$ | $\begin{gathered} 0.040 \\ (0.002) \end{gathered}$ | $\begin{gathered} 0.037 \\ (0.005) \end{gathered}$ | $\begin{gathered} 0.036 \\ (0.002) \end{gathered}$ | $\begin{gathered} 0.041 \\ (0.002) \end{gathered}$ | $\begin{gathered} 0.031 \\ (0.002) \end{gathered}$ | $\begin{gathered} 0.044 \\ (0.008) \end{gathered}$ | $\begin{gathered} 0.031 \\ (0.005) \end{gathered}$ | $\begin{gathered} 0.043 \\ (0.008) \end{gathered}$ |
| Num. of obs. | 1,553,021 | 1,287,773 | 687,323 | 1,553,021 | 1,553,021 | 1,473,218 | 1,553,021 | 1,553,021 | 1,553,021 |

## Section 2: Neighborhood Exposure Effects in Denmark

## Data

■ Data source: Danish registers
■ Data span: 1982-2000
■ Sample: Children who were born between 1970-1982

- permanent residents (stayers/PR): subset of parents who reside in a single municipality (parish) c in 1982-2000
- movers: individuals in the main sample who are not PR

■ Income type: Disposable income

- averaged over 1982-2000 to get parental income

■ Unit of Analysis: Family income for parents and individual income for children

## Summary Statistics

## Population Distribution

Figure: Distribution of Population across Municipalities


## Education Level and PR Status

Figure: Distribution of Years of Schooling by Permanent Residence Status


## Characteristics of Movers and Stayers

Figure: Income Distribution of Parents: Movers vs Non-movers
(a) Individual Level

(b) Household Level


## Number of Moves

Figure: Number of Moves by Education Level


# Neighborhood Exposure Effects 

## Mean Income Ranks for Children of PR of Copenhagen

Figure: Mean Child Inc. Rank vs Parent Inc. Rank for Children


## Movers' Outcomes versus Predicted Outcomes Based on PR in Destination- Movers at Age 13



## Childhood Exposure Effect Estimates- Specification

$$
\begin{aligned}
y_{i}= & \sum_{s=1980}^{1988} I\left(s_{i}=s\right)\left(\alpha_{s}^{1}+\alpha_{s}^{2} \bar{y}_{p o s}\right)+\sum_{m=9}^{30} I\left(m_{i}=m\right)\left(\zeta_{m}^{1}+\zeta_{m}^{2} p_{i}\right) \\
& +\sum_{s=1980}^{1987} \kappa_{s}^{d} I\left(s_{i}=s\right) \Delta_{o d p s}+I\left(m_{i} \leqslant 23\right)\left(b_{0}+\left(23-m_{i}\right) \gamma\right) \Delta_{o d p s} \\
(7) \quad & +I\left(m_{i}>23\right)\left(\delta+\left(23-m_{i}\right) \delta^{\prime}\right) \Delta_{o d p s}+\varepsilon_{3 i}
\end{aligned}
$$

## Childhood Exposure Effects on Inc. Ranks

Figure: Childhood Exposure Effects on Income Ranks in Adulthood


## Childhood Exposure Effect Estimates

Table 1: Childhood Exposure Effect Estimates
Dependent variable: Child's income rank at age 26 (30)

|  |  |  |  |  |  |  | Family FE |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Spec: | Pooled <br> (1) | $\text { Age }<=23$ <br> (2) | $\text { Age }<18$ <br> (3) | No cohort controls <br> (4) | Family Income (5) | Child nbhd FE <br> (6) | Baseline <br> (7) | No cohort controls (8) | Timevarying controls (9) |
| US ( $\gamma$ ) | $\begin{gathered} 0.040 \\ (0.002) \end{gathered}$ | $\begin{gathered} 0.040 \\ (0.002) \end{gathered}$ | $\begin{gathered} 0.037 \\ (0.005) \end{gathered}$ | $\begin{gathered} 0.036 \\ (0.002) \end{gathered}$ | $\begin{gathered} 0.041 \\ (0.002) \end{gathered}$ | $\begin{gathered} 0.031 \\ (0.002) \end{gathered}$ | $\begin{gathered} 0.044 \\ (0.008) \end{gathered}$ | $\begin{gathered} 0.031 \\ (0.005) \end{gathered}$ | $\begin{gathered} 0.043 \\ (0.008) \end{gathered}$ |
| DK : | $\begin{gathered} 0.013 \\ (0.003) \\ \hline \end{gathered}$ | $\begin{gathered} 0.014 \\ (0.003) \\ \hline \end{gathered}$ | $\begin{array}{r} 0.015 \\ (0.004) \\ \hline \end{array}$ | $\begin{array}{r} 0.011 \\ (0.003) \\ \hline \end{array}$ | $\begin{gathered} 0.014 \\ (0.003) \\ \hline \end{gathered}$ | $\begin{gathered} 0.012 \\ (0.003) \\ \hline \end{gathered}$ | $\begin{array}{r} 0.009 \\ (0.009) \\ \hline \end{array}$ | $\begin{gathered} 0.012 \\ (0.008) \\ \hline \end{gathered}$ | $\begin{array}{r} -0.008 \\ (0.016) \\ \hline \end{array}$ |

# Section 3: Discussion 

# (A) Selection and Age of Child at Move: 

(A.i) Parental Characteristics

## (A.i-1) Education

## Parental Char. and Age of Child when Parents Move

Figure: Age of Child at Move and Parental Edu. by Ownership Status
(a) Owners
(b) Renters



## Distribution of Age of Child at Move- by Ownership Status

Figure: Timing of Moves across Neighborhoods by Home Ownership


# (A) Selection and Age of Child at Move: 

(A.i) Parental Characteristics

## (A.i-2) Disposable Income

## Parental Disposable Inc. by Ownership Status

Figure: Parental Income Rank and Age of Child when Parents Move
(a) Owners
(b) Renters



# (A) Selection and Age of Child at Move: 

(A.i) Parental Characteristics

## (A.i-3) Family Structure

## Parental Family Structure and Age of Child at Move

Figure: Fraction of Intact Families and Age of Child when Parents Move
(a) Owners

(b) Renters


# (A) Selection and Age of Child at Move: 

(A.i) Parental Characteristics

## (A.i-4) Family Size

## Family Size and Age of Child at Move

Figure: Average Family Size and Age of Child when Parents Move
(a) Owners

(b) Renters

(B) Parental Sorting to Neighborhoods:
(B.i) Quality of Moves

# (B.i-1) Predicted Difference in Predicted Outcomes of Children in Orig. vs Dest. 

## The Quality of Moves and Age of Child at Move

Figure: The Quality of Moves by Ownership Status
(a) Owners
(b) Renters



## The Quality of Origin and Age of Child at Move

Figure: The Quality of Moves by Ownership Status
(a) Owners
(b) Renters



## The Quality of Destination and Age of Child at Move

Figure: The Quality of Moves by Ownership Status
(a) Owners
(b) Renters


(B) Parental Sorting to Neighborhoods:
(B.i) Quality of Moves
(B.i-2) NBHD Avg Inc Rank at Orig. vs Dest.

## NBHD (Parish) Income Rank and Age of Child at Move

Figure: Change in nbhd Inc Rank and Age of Child
(a) Owners
(b) Renters



## NBHD (Large Clusters) Inc Rank and Age of Child at Move

Figure: Change in nbhd Inc Rank and Age of Child
(a) Owners
(b) Renters



## NBHD (Small Clusters) Inc Rank and Age of Child at Move

Figure: Change in nbhd Inc Rank and Age of Child
(a) Owners
(b) Renters


(B) Parental Sorting to Neighborhoods:
(B.i) Quality of Moves
(B.i-3) School Quality Rank at Orig. vs Dest.

## NBHD (Parish) School Quality Rank and Age at Move

Figure: Change in nbhd School Rank (Math Grades) and Age of Child
(a) Owners

(b) Renters

(B.i-4) Average Neighborhood House Price Rank at 200-HH Block Level

## Neighborhood House Price Rank and Age of Child at Move

Figure: Change in nbhd House Price Rank. and Age of Child
(a) Owners
(b) Renters


(C) Timing of Moves and Lifecycle Shocks

## (C.i) Divorce

## Divorce and Age of Child at Move

Figure: Age of Child at Move \& Frac. of Parents Separated when Moving
(a) Owners
(b) Renters


(C) Timing of Moves and Lifecycle Shocks

# (C.ii) Change to Income when Moving 

Figure: Age of Child at Move and the Change to Family Disp. Inc. Rank
(a) Owners
(b) Renters



Figure: Age of Child at Move and the Change to Family Wage Inc. Rank
(a) Owners
(b) Renters



## (D) Family Fixed Effect and Exogeneity Assumption

## Family Fixed Effect Model

- Authors address time-varying selection possibility by adding family FE to the parametric model (and, separately, by controlling for changes in parents' income and marital status):

$$
y_{i}=\sum_{s=1980}^{1988} I\left(s_{i}=s\right)\left(\alpha_{s}^{1}+\alpha_{s}^{2} \bar{y}_{p o s}\right)+\sum_{m=9}^{30} I\left(m_{i}=m\right)\left(\zeta_{m}^{1}+\zeta_{m}^{2} p_{i}\right)
$$

$$
\begin{equation*}
+\sum_{m=9}^{30} b_{m} I\left(m_{i}=m\right) \Delta_{o d p s}+\sum_{s=1980}^{1987} \kappa_{s}^{d} I\left(s_{i}=s\right) \Delta_{o d p s}+\varepsilon_{3 i} . \tag{6}
\end{equation*}
$$

- Regression is now should estimated entirely on sample of families with 2 children. Intuitively, family-level mean effects are taken out.


## Childhood Exposure Effect Estimates- Results

TABLE II
Childhood Exposure Effect Estimates

| Specification: | Dependent variable: Child's income rank at age 24 |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  | With | amily fixed | ffects |
|  | Pooled <br> (1) | $\text { Age } \leqslant 23$ <br> (2) | $\text { Age }<18$ <br> (3) | No cohort controls <br> (4) | Individual income <br> (5) | Child CZ FE <br> (6) | Baseline <br> (7) | No cohort controls (8) | Timevarying controls (9) |
| Exposure effect ( $\gamma$ ) | $\begin{gathered} 0.040 \\ (0.002) \end{gathered}$ | $\begin{gathered} 0.040 \\ (0.002) \end{gathered}$ | $\begin{gathered} 0.037 \\ (0.005) \end{gathered}$ | $\begin{gathered} 0.036 \\ (0.002) \end{gathered}$ | $\begin{gathered} 0.041 \\ (0.002) \end{gathered}$ | $\begin{gathered} 0.031 \\ (0.002) \end{gathered}$ | $\begin{gathered} 0.044 \\ (0.008) \end{gathered}$ | $\begin{gathered} 0.031 \\ (0.005) \end{gathered}$ | $\begin{gathered} 0.043 \\ (0.008) \end{gathered}$ |
| Num. of obs. | 1,553,021 | 1,287,773 | 687,323 | 1,553,021 | 1,553,021 | 1,473,218 | 1,553,021 | 1,553,021 | 1,553,021 |

## Discussion: Family Fixed Effect Model

- Suppose we can write $\epsilon_{i}=\hat{\theta}_{\text {fam }, i}+e_{i}$
- $\hat{\theta}_{\text {fam, }, \text { : }}$ fixed family inputs (culture, parents' HC, etc.)
- $e_{i}$ : variable inputs (e.g., wealth shocks, noise)
- The selection assumption: $\delta_{m}=\frac{\operatorname{cov}\left(\epsilon_{i}, \bar{y}_{p d s}\right)}{\operatorname{var}\left(\bar{y}_{p d s}\right)}$ is constant in age
- Including family fixed effects controls for $\hat{\theta}_{\text {fam }}$ : if higher-skill families choose better neighborhoods at earlier ages
- To interpret results as causal still need $\frac{\operatorname{cov}\left(e_{i}, \bar{y}_{p d s}\right)}{\operatorname{var}\left(\bar{y}_{p d s}\right)}$ cons. in age
- May be violated if shocks to wealth are corr. with child's age
- One such shock correlated with first child's age: the birth of a $2^{\text {nd }}$ child
- Meaningful differences between families where kids are 2 years vs. 8 years apart.

Figure: Time Space and Differences in Sibling Outcomes


Figure: Time Space and Choices of Destination Neighborhood


## Discussion of the Identification Assumption

## Identification Assumption

Exposure effect at age $m$ : the impact of spending year $m$ of one's childhood in an area where PR's outcomes are 1 pp higher

Thought experiment: randomly assign children to new nbhd $d$ starting at age $m$ for the rest of childhood. The best linear predictor of children's outcomes $y_{i}$ in the experimental sample, based on the PR's outcomes in CZ $d\left(\bar{y}_{p d s}\right)$ :

$$
\begin{equation*}
y_{i}=\alpha_{m}+\beta_{m} \bar{y}_{p d s}+\theta_{i} \tag{3}
\end{equation*}
$$

Random assignment: $\theta \perp \bar{y}_{p d s}$
Exposure effect at $m: \gamma_{m}=\beta_{m}-\beta_{m+1}$, the effect on $y_{i}$ of spending the year from age m to age $\mathrm{m}+1$ in the destination Observational data: $b_{m}=\beta_{m}+\delta_{m}$
Bias $=\delta_{m}=\frac{\operatorname{cov}\left(\theta_{i}, \bar{y}_{p d s}\right)}{\operatorname{var}\left(\bar{y}_{p d s}\right)}$ : parent inputs \& unobserved det. of children's outcomes covary with PR's outcomes

## Exposure Effects- Constant-in-Age Selection Assumption

Bias $=\delta_{m}=\frac{\operatorname{cov}\left(\theta_{i}, \bar{y}_{p d s}\right)}{\operatorname{var}\left(\bar{y}_{p d s}\right)}$
ASSUMPTION 1 (A.1): Selection effects do not vary with the child's age at move: $\delta_{m}=\delta$ for all m.

Under A.1, we obtain consistent estimates of exposure effects:

$$
\gamma_{m}=\left(\beta_{m}+\delta_{m}\right)-\left(\beta_{m+1}+\delta_{m+1}\right)=b_{m}-b_{m+1}
$$

Even in observational data because the selection effects $\delta$ cancel out when estimating the exposure effect.

Rules out differential preferences among parents by age of child for local amenities, such as school quality, that are not fully captured in adult income percentile rank $\bar{y}_{p d s}$

## What if Assumption A. 1 Is violated?

Under A.1:

$$
\gamma_{m}=\left(\beta_{m}-\beta_{m+1}\right)+\left(\delta_{m}-\delta_{m+1}\right)=b_{m}-b_{m+1}
$$

If A. 1 is violated:
1 If sorting decreases in child's age:
$\delta_{m}>\delta_{m+1} \quad \forall m \in\{\underline{m}, \ldots, \bar{m}\} \Rightarrow$ equ (3) overestimates the exposure effect, $\gamma_{m}$

2 If sorting becomes stronger as age increases:
$\delta_{m}<\delta_{m+1} \quad \forall m \in\{\underline{m}, \ldots, \bar{m}\} \Rightarrow$ equ (3) underestimates the exposure effect, $\gamma_{m}$.
3 Unclear if sorting not monotonically changes over the age support exploited for the estimation.

## Parental Selection based on Education

Chetty (2018) estimates:

$$
\begin{equation*}
y_{i}=\alpha+\beta_{m} \Delta_{o d p s}+\epsilon_{i} \tag{4}
\end{equation*}
$$

Parent's education level is one of the omitted variables affecting both child's outcome and quality of the move across NBHDs.

Let's assume that the true model is as follows:

$$
\begin{equation*}
y_{i}=\alpha+\beta_{m} \Delta_{o d p s}+\beta_{e} e d u_{i}^{p}+u_{i} \tag{5}
\end{equation*}
$$

Then,

$$
\begin{aligned}
& \operatorname{Plim} \hat{\beta}_{m}= \beta_{m}+\beta_{e} \frac{\operatorname{cov}\left(e d u_{i}^{p}, \Delta_{p d s}\right)}{\operatorname{var}\left(\Delta_{p d s}\right)} \\
&=\beta_{m}+\beta_{e} \delta_{m}
\end{aligned}
$$

$$
\text { Plim } \hat{\gamma}_{m}=\left(\beta_{m}-\beta_{m+1}\right)+\beta_{e}\left(\delta_{m}-\delta_{m+1}\right)
$$

## Intensity of Sorting by Age of Child at Move

Figure: Intensity of Sorting b/w Parent's Education and Quality of Move
(a) Owners
(b) Renters



## Intensity of Sorting by Age of Child at Move

Figure: Intensity of Sorting b/w Family Structure and Quality of Move
(a) Owners
(b) Renters



## Intensity of Sorting by Age of Child at Move

Figure: Intensity of Sorting b/w Ownership Status and Quality of Move


## Back-of-the-envelope Calculation of the Bias

To evaluate the size of the bias, $\beta_{e}\left(\delta_{m}-\delta_{m+1}\right)$ :
1 Using equ (5), obtain some estimates for $\beta_{e}: \hat{\beta}_{e} \in[0.82,1.15]$
2 Using the slope of covariance term (between parents' education level and quality of the move) over age of child, obtain an estimate for $\left(\delta_{m}-\delta_{m+1}\right):\left(\delta_{m}-\delta_{m+1}\right) \approx 0.005$

## Conclusion

- Recent studies have exploited quasi-experimental strategies to identify the causal impact of NBHDs to outcomes of children.
- One of the main challenges in estimating the causal impact of NBHDs on child is the endogeneity of NBHD quality.
- We investigated the main identifying assumptions of recent studies in the literature.

■ Parental sorting into NBHDs has an important lifecycle gradient; it is not orthogonal to age of children at move.

- The constant-in-age selection effects assumption in recent empirical works is violated, leading researchers to overestimate the impact of NBHD on child outcomes.


## Thanks!

## Appendix

## Geographical Variation in Outcomes of PR- across CZs



## Mean Income Ranks for Children of PR of Copenhagen

Figure: Mean Child Inc. Rank vs Parent Inc. Rank for Children


## Population Distribution

Figure: Distribution of Population (Parish-level)


## Education Level and PR Status

Figure: Distribution of Years of Schooling by Permanent Residence Status

$\square$

## Number of Moves

Figure: Number of Moves by Education Level


## Parental Char. and Age of Child when Parents Move

Figure: Age of Child at Move and Parental Edu. by Ownership Status
(a) Owners
(b) Renters



## Distribution of Age of Child at Move- by Ownership Status

Figure: Timing of Moves across Neighborhoods by Home Ownership


## Parental Disposable Inc. by Ownership Status

Figure: Parental Income Rank and Age of Child when Parents Move
(a) Owners
(b) Renters



## Parental Family Structure and Age of Child at Move

Figure: Fraction of Intact Families and Age of Child when Parents Move
(a) Owners

(b) Renters


## Family Size and Age of Child at Move

Figure: Average Family Size and Age of Child when Parents Move
(a) Owners

(b) Renters


## The Quality of Moves and Age of Child at Move

Figure: The Quality of Moves by Ownership Status
(a) Owners

(b) Renters


## The Quality of Origin and Age of Child at Move

Figure: The Quality of Moves by Ownership Status
(a) Owners
(b) Renters



## The Quality of Destination and Age of Child at Move

Figure: The Quality of Moves by Ownership Status
(a) Owners
(b) Renters



## NBHD (Parish) Income Rank and Age of Child at Move

Figure: Change in nbhd Inc Rank and Age of Child
(a) Owners

(b) Renters


## NBHD (Large Clusters) Inc Rank and Age of Child at Move

Figure: Change in nbhd Inc Rank and Age of Child
(a) Owners
(b) Renters



## NBHD (Small Clusters) Inc Rank and Age of Child at Move

Figure: Change in nbhd Inc Rank and Age of Child
(a) Owners
(b) Renters



## NBHD (Parish) School Quality Rank and Age at Move

Figure: Change in nbhd School Rank (Math Grades) and Age of Child
(a) Owners

(b) Renters


## Neighborhood House Price Rank and Age of Child at Move

Figure: Change in nbhd House Price Rank. and Age of Child
(a) Owners
(b) Renters



## Divorce and Age of Child at Move

Figure: Age of Child at Move \& Frac. of Parents Separated when Moving
(a) Owners



