

Factor Models: A Review

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Factor Models: A Review of General Models

$$E(\theta) = 0; \quad E(\varepsilon_i) = 0; \quad i = 1, \dots, 5$$
$$\theta \perp\!\!\!\perp (\varepsilon_1, \dots, \varepsilon_5)$$

$$R_1 = \alpha_1\theta + \varepsilon_1, \quad R_2 = \alpha_2\theta + \varepsilon_2, \quad R_3 = \alpha_3\theta + \varepsilon_3,$$
$$R_4 = \alpha_4\theta + \varepsilon_4, \quad R_5 = \alpha_5\theta + \varepsilon_5, \quad \varepsilon_i \perp\!\!\!\perp \varepsilon_j, \quad i \neq j$$

$$\text{Cov}(R_1, R_2) = \alpha_1\alpha_2\sigma_\theta^2$$

$$\text{Cov}(R_1, R_3) = \alpha_1\alpha_3\sigma_\theta^2$$

$$\text{Cov}(R_2, R_3) = \alpha_2\alpha_3\sigma_\theta^2$$

- Normalize $\alpha_1 = 1$

$$\frac{\text{Cov}(R_2, R_3)}{\text{Cov}(R_1, R_2)} = \alpha_3$$

- \therefore We know σ_θ^2 from $\text{Cov}(R_1, R_2)$.
- From $\text{Cov}(R_1, R_3)$ we know

$$\alpha_3, \alpha_4, \alpha_5.$$

- Can get the variances of the ε_i from variances of the R_i

$$\text{Var}(R_i) = \alpha_i^2 \sigma_\theta^2 + \sigma_{\varepsilon_i}^2.$$

- If $T = 2$, all we can identify is $\alpha_1 \alpha_2 \sigma_\theta^2$.
- If $\alpha_1 = 1$, $\sigma_\theta^2 = 1$, we identify α_2 .
- Otherwise model is fundamentally underidentified.

2 Factors: (Some Examples)

$$\theta_1 \perp\!\!\!\perp \theta_2$$

$$\varepsilon_i \perp\!\!\!\perp \varepsilon_j \quad \forall i \neq j$$

$$R_1 = \alpha_{11}\theta_1 + (0)\theta_2 + \varepsilon_1$$

$$R_2 = \alpha_{21}\theta_1 + (0)\theta_2 + \varepsilon_2$$

$$R_3 = \alpha_{31}\theta_1 + \alpha_{32}\theta_2 + \varepsilon_3$$

$$R_4 = \alpha_{41}\theta_1 + \alpha_{42}\theta_2 + \varepsilon_4$$

$$R_5 = \alpha_{51}\theta_1 + \alpha_{52}\theta_2 + \varepsilon_5$$

Let $\alpha_{11} = 1$, $\alpha_{32} = 1$. (Set scale)

$$\text{Cov}(R_1, R_2) = \alpha_{21}\sigma_{\theta_1}^2$$

$$\text{Cov}(R_1, R_3) = \alpha_{31}\sigma_{\theta_1}^2$$

$$\text{Cov}(R_2, R_3) = \alpha_{21}\alpha_{31}\sigma_{\theta_1}^2$$

- Form ratio of $\frac{\text{Cov}(R_2, R_3)}{\text{Cov}(R_1, R_2)} = \alpha_{31}$, \therefore we identify $\alpha_{31}, \alpha_{21}, \sigma_{\theta_1}^2$, as before.

$$\text{Cov}(R_1, R_4) = \alpha_{41}\sigma_{\theta_1}^2, \quad \therefore \text{since we know } \sigma_{\theta_1}^2 \therefore \text{we get } \alpha_{41}.$$

\vdots

$$\text{Cov}(R_1, R_k) = \alpha_{k1}\sigma_{\theta_1}^2$$

- \therefore we identify α_{k1} for all k and $\sigma_{\theta_1}^2$.

$$\begin{aligned} \text{Cov}(R_3, R_4) - \alpha_{31}\alpha_{41}\sigma_{\theta_1}^2 &= \alpha_{42}\sigma_{\theta_2}^2 \\ \text{Cov}(R_3, R_5) - \alpha_{31}\alpha_{51}\sigma_{\theta_1}^2 &= \alpha_{52}\sigma_{\theta_2}^2 \\ \text{Cov}(R_4, R_5) - \alpha_{41}\alpha_{51}\sigma_{\theta_1}^2 &= \alpha_{52}\alpha_{42}\sigma_{\theta_2}^2, \end{aligned}$$

- By same logic,

$$\frac{\text{Cov}(R_4, R_5) - \alpha_{41}\alpha_{51}\sigma_{\theta_1}^2}{\text{Cov}(R_3, R_4) - \alpha_{31}\alpha_{41}\sigma_{\theta_1}^2} = \alpha_{52}$$

- \therefore get $\sigma_{\theta_2}^2$ of “2” loadings.

- If we have dedicated measurements on each factor do not need a normalization on the factors of R .
- Dedicated measurements set the scales and make factor models interpretable:

$$M_1 = \theta_1 + \varepsilon_{1M}$$

$$M_2 = \theta_2 + \varepsilon_{2M}$$

$$\text{Cov}(R_1, M) = \alpha_{11}\sigma_{\theta_1}^2$$

$$\text{Cov}(R_2, M) = \alpha_{21}\sigma_{\theta_1}^2$$

$$\text{Cov}(R_3, M) = \alpha_{31}\sigma_{\theta_1}^2$$

$$\text{Cov}(R_1, R_2) = \alpha_{11}\alpha_{12}\sigma_{\theta_1}^2,$$

$$\text{Cov}(R_1, R_3) = \alpha_{11}\alpha_{13}\sigma_{\theta_1}^2, \quad \text{so we can identify } \alpha_{12}\sigma_{\theta_1}^2$$

- \therefore We can get $\alpha_{12}, \sigma_{\theta_1}^2$ and the other parameters.

General Case

$$R_{T \times 1} = M_{T \times 1} + \Lambda_{T \times K} \theta_{K \times 1} + \varepsilon_{T \times 1}$$

- θ are factors, ε uniquenesses, $\theta \perp\!\!\!\perp \varepsilon$

$$E(\varepsilon) = 0$$
$$\text{Var}(\varepsilon\varepsilon') = D = \begin{pmatrix} \sigma_{\varepsilon_1}^2 & 0 & \cdots & 0 \\ 0 & \sigma_{\varepsilon_2}^2 & 0 & \vdots \\ \vdots & 0 & \ddots & \vdots \\ 0 & \cdots & 0 & \sigma_{\varepsilon_T}^2 \end{pmatrix}$$

$$E(\theta) = 0$$
$$\text{Var}(R) = \Lambda \Sigma_{\theta} \Lambda' + D \quad \Sigma_{\theta} = E(\theta\theta')$$

- The only source of information on Λ and Σ_θ is from the covariances.
- (Each variance is “contaminated” by a uniqueness.)
- Associated with each variance of R_i is a $\sigma_{\varepsilon_i}^2$.
- Each uniqueness variance contributes one new parameter.
- How many unique covariance terms do we have?
$$\frac{T(T-1)}{2}.$$

- We have T uniquenesses; TK elements of Λ .
- $\frac{K(K-1)}{2}$ elements of Σ_{θ} .
- $\frac{K(K-1)}{2} + TK$ parameters $(\Sigma_{\theta}, \Lambda)$.
- Need this many covariances to identify model
 “Ledermann Bound”:

$$\frac{T(T-1)}{2} \geq TK + \frac{K(K-1)}{2}$$

Lack of Identification Up to Rotation

- Observe that if we multiply Λ by an orthogonal matrix C , ($CC' = I$), we obtain

$$\text{Var}(R) = \Lambda C [C' \Sigma_{\theta} C] C' \Lambda' + D$$

- C is a “rotation.”
- Cannot separate ΛC from Λ .
- Model not identified against orthogonal transformations in the general case.

Some common assumptions:

① $\theta_i \perp\!\!\!\perp \theta_j, \forall i \neq j$

$$\Sigma_{\theta} = \begin{pmatrix} \sigma_{\theta_1}^2 & 0 & \cdots & 0 \\ 0 & \sigma_{\theta_2}^2 & 0 & \vdots \\ \vdots & 0 & \ddots & \vdots \\ 0 & \cdots & 0 & \sigma_{\theta_K}^2 \end{pmatrix}$$

joined with

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$$\Lambda = \begin{pmatrix} 1 & 0 & 0 & 0 & \cdots & 0 \\ \alpha_{21} & 0 & 0 & 0 & \cdots & 0 \\ \alpha_{31} & 1 & 0 & 0 & \cdots & 0 \\ \alpha_{41} & \alpha_{42} & 0 & 0 & \cdots & 0 \\ \alpha_{51} & \alpha_{52} & 1 & 0 & \cdots & 0 \\ \alpha_{61} & \alpha_{62} & \alpha_{63} & 0 & \cdots & 0 \\ \vdots & \vdots & \vdots & 1 & & \vdots \end{pmatrix}$$

- We know that we can identify of the Λ, Σ_{θ} parameters.

$$\frac{K(K-1)}{2} + TK \leq \frac{T(T-1)}{2}$$

of free parameters data
"Ledermann Bound"

- Can get more information by looking at higher order moments.
- (See, e.g., Bonhomme and Robin, 2009.)

- Normalize: $\alpha_{I^*} = 1, \alpha_1 = 1 \quad \therefore \sigma_\theta^2 \quad \therefore \alpha_1$.
- Can make alternative normalizations.

Recovering the Distributions Nonparametrically

Theorem 1

Suppose that we have two random variables T_1 and T_2 that satisfy:

$$T_1 = \theta + v_1$$

$$T_2 = \theta + v_2$$

with θ, v_1, v_2 mutually statistically independent, $E(\theta) < \infty$, $E(v_1) = E(v_2) = 0$, that the conditions for Fubini's theorem are satisfied for each random variable, and the random variables possess nonvanishing (a.e.) characteristic functions, then the densities $f(\theta)$, $f(v_1)$, and $f(v_2)$ are identified.

Proof.

See Kotlarski (1967). □

- Suppose

$$\begin{aligned}I &= \mu_I(X, Z) + \alpha_I\theta + \varepsilon_I \\Y_0 &= \mu_0(X) + \alpha_0\theta + \varepsilon_0 \\Y_1 &= \mu_1(X) + \alpha_1\theta + \varepsilon_1 \\M &= \mu_M(X) + \theta + \varepsilon_M.\end{aligned}$$

- System can be rewritten as

$$\begin{aligned}\frac{I - \mu_I(X, Z)}{\alpha_I} &= \theta + \frac{\varepsilon_I}{\alpha_I} \\ \frac{Y_0 - \mu_0(X)}{\alpha_0} &= \theta + \frac{\varepsilon_0}{\alpha_0} \\ \frac{Y_1 - \mu_1(X)}{\alpha_1} &= \theta + \frac{\varepsilon_1}{\alpha_1} \\ M - \mu_M(X) &= \theta + \varepsilon_M\end{aligned}$$

- Applying Kotlarski's theorem, identify the densities of

$$\theta, \frac{\varepsilon_I}{\alpha_I}, \frac{\varepsilon_0}{\alpha_0}, \frac{\varepsilon_1}{\alpha_1}, \varepsilon_M.$$

- We know α_I , α_0 and α_1 .
- Can identify the densities of $\theta, \varepsilon_I, \varepsilon_0, \varepsilon_1, \varepsilon_M$.
- Recover the joint distribution of (Y_1, Y_0) .

$$F(Y_1, Y_0 | X) = \int F(Y_1, Y_0 | \theta, X) dF(\theta).$$

- $F(\theta)$ is known.

$$F(Y_1, Y_0 | \theta, X) = F(Y_1 | \theta, X) F(Y_0 | \theta, X).$$

- $F(Y_1 | \theta, X)$ and $F(Y_0 | \theta, X)$ identified

$$F(Y_1 | \theta, X, S = 1) = F(Y_1 | \theta, X)$$

$$F(Y_0 | \theta, X, S = 0) = F(Y_0 | \theta, X).$$

- Can identify the number of factors generating dependence among the Y_1, Y_0, C, S and M .