# Tasks and Heterogeneous Human Capital

## Shintaro Yamaguchi. (2012). Journal of Labor Economics, 30(1): pp. 1–53.

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Tasks and Heterogeneous Human Capital

## I. Introduction



#### II. Model



- Task complexity: x<sub>t</sub>
- Skills: *s*<sub>t</sub>

## Skills vs. Task Complexity



- Jobs in the same occupation are homogeneous in terms of task complexity; that is, jobs and occupations are not distinguished.
- In each year t, an individual chooses an occupation that lies in a K-dimensional continuous space of task complexity x<sub>t</sub> that is observable and takes nonnegative values.
- The **task complexity** indexes take nonnegative values, and sufficiently many occupations exist so that an individual can choose any occupation in the task complexity space.
- **Skills** in year *t* are denoted by a *K*-dimensional vector *s*<sub>t</sub> that is unobserved by the econometrician.
- The skills index s<sub>t</sub> can take any real number, including negative values.



- Labor is the only factor of production.
- Each firm offers jobs of a single type of complexity, which implies that the products of each firm can be characterized by a task complexity vector.
- The products are heterogeneous and consumed by households.
- The price of the product characterized by task  $x_t$  is denoted by  $\pi(x_t)$ .
- The productivity of a worker with skill  $s_t$  in a job with task complexity  $x_t$  is  $q(x_t, s_t)$ .



## A. Wage Function



• The marginal value product of a worker with skill s<sub>t</sub> in an occupation with task complexity x<sub>t</sub> is

$$w_t = \pi(x_t)q(x_t, s_t)\exp(\eta_t), \qquad (1)$$

- $\eta_t \sim N(0, \sigma_{\eta}^2)$  can be interpreted as an independent and identically distributed productivity shock or as measurement error.
- As in the Roy model, skills are rewarded differently across occupations.
- (1) is assumed.



• Labor productivity:

$$\ln q(x_t, s_t) = \theta'(x_t)s_t, \qquad (2)$$

- θ(x<sub>t</sub>) is a K-dimensional vector of implicit skill prices and represents the contribution of skills s<sub>t</sub> to an occupation with task x<sub>t</sub>.
- Skills are more intensely used and contribute to productivity more, when the corresponding tasks are complex ∂θ<sub>k</sub>(x)/∂x<sub>k</sub> > 0, where subscript k is an index for the task dimension.



I parameterize the output price and productivity:

$$\ln \pi(x_t) = p_0 + p_1' x_t$$
 (3)

and

$$\ln q(x_t, s_t) = \theta'(x_t)s_t = [p_2 + P'_3 x_t]'s_t,$$
(4)

- *p*<sub>0</sub> is a scalar,
- *p*<sub>1</sub> and *p*<sub>2</sub> are *K*-dimensional vectors, and
- P<sub>3</sub> is a K-dimensional diagonal matrix.
- Therefore, the log wage is given by

$$\ln w_t = p_0 + p_1' x_t + [p_2 + P_3' x_t]' s_t + \eta_t.$$
(5)



#### **B. Skill Formation**



- Let *d* be an *L*-dimensional vector of individual characteristics that are fixed at labor market entry, such as race and education.
- A vector of skill shocks ε<sub>t</sub> is normal, independent, and identically distributed with mean zero and variance Σ<sub>ε</sub> : ε<sub>t</sub> ~ N(0, Σ<sub>ε</sub>).
- Skills grow from year t to year t + 1 according to the following skill transition equation:

$$s_{t+1} = Ds_t + a_0 + A_1x_t + A_2d + \varepsilon_{t+1}, \qquad (6)$$

- D is a K-dimensional diagonal matrix for skill depreciation,
- *a*<sup>0</sup> is a *K*-dimensional vector of parameters,
- A<sub>1</sub> is a  $K \times K$  diagonal matrix of the marginal effects of task complexity on learning, and
- A<sub>2</sub> is a K × L-dimensional matrix that represents heterogeneous learning ability.

- Individuals start their careers with initial skills s<sub>1</sub> that differ across individuals in both observable and unobservable ways.
- The initial skill endowment:

$$s_1 = h + Hd + \varepsilon_1, \tag{7}$$

- h is a K-dimensional vector,
- *H* is a  $K \times L$  matrix of parameters,
- *d* is a vector of observed individual characteristics at labor market entry, and
- ε<sub>1</sub> is an unobserved component of initial skills that is distributed as ε<sub>1</sub> ~ N(0, Σ<sub>ε1</sub>).



#### C. Job Preferences



• Utility derived from work:

$$v_{t} = v(x_{t}, \bar{x}_{t}, s_{t}, \tilde{\nu}_{t}; d)$$

$$= (g_{0} + G_{1}d + G_{2}s_{t} + \tilde{\nu}_{t})'x_{t} + x_{t}'G_{3}x_{t} + (x_{t} - \bar{x}_{t})'G_{4}(x_{t} - \bar{x}_{t}),$$
(9)

- g<sub>0</sub> is a K-dimensional vector of preference parameters;
- $G_1$  is a  $K \times L$  matrix of preference parameters;
- $\tilde{\nu}_t$  is a K-dimensional vector of preference shocks with zero mean;
- $G_2$ ,  $G_3$ , and  $G_4$  are  $K \times K$  diagonal matrices; and
- $\bar{x}_t$  is a *K*-dimensional vector of work habits.



- The utility from job tasks varies across individuals according to their individual characteristics *d*, skill levels *s*<sub>t</sub>, a preference shock *ν*<sub>t</sub>, and work habits *x*<sub>t</sub>.
- Skilled workers prefer complex tasks if the parameter matrix *G*<sub>2</sub> is positive definite.
- *G*<sub>3</sub>: negative definite.
- For a very high value of x<sub>t</sub>, the marginal utility from task complexity is negative; this is the cost of entering an occupation with complex tasks.
- The parameters  $g_0$ ,  $G_1$ , and  $G_2$  are unrestricted.
- The last term in the above equation captures the effect of work habits on utility.



Individuals form their work habits x
<sub>t+1</sub> through the following transition equation:

$$\bar{x}_{t+1} = A_3 \bar{x}_t + (I - A_3) x_t,$$
 (10)

- A<sub>3</sub> is a *K*-dimensional diagonal matrix of which elements take values between zero and one, and
- *I* is a *K*-dimensional identity matrix.



• The initial condition for  $\bar{x}_t$  varies across individuals according to initial observed characteristics *d* such that

$$\bar{x}_1 = \bar{x}_{1,0} + Xd,$$
 (11)

- where  $\bar{x}_{1,0}$  is a K-dimensional vector of parameters, and
- X is a  $K \times L$  matrix of parameters.



### **D. Bellman Equation**



The Bellman equation for an individual is given by

$$V_{t}(s_{t}, \bar{x}_{t}, \tilde{\nu}_{t}, \eta_{t}; d) = \max_{x_{t}} \ln w(x_{t}, s_{t}, \eta_{t}) + v(x_{t}, \bar{x}_{t}, s_{t}, \tilde{\nu}_{t}; d) + \beta E V_{t+1}(s_{t+1}, \bar{x}_{t+1}, \tilde{\nu}_{t+1}, \eta_{t+1}; d),$$
(12)

$$\ln w_t = p_0 + p_1' x_t + [p_2 + P_3' x_t]' s_t + \eta_t,$$
(13)

$$v_t = (g_0 + G_1 d + G_2 s_t + \tilde{\nu}_t)' x_t + x'_t G_3 x_t + (x_t - \bar{x}_t)' G_4 (x_t - \bar{x}_t), \qquad (14)$$

$$s_{t+1} = Ds_t + a_0 + A_1 x_t + A_2 d + \varepsilon_{t+1}, \qquad (15)$$

$$\bar{x}_{t+1} = A_3 \bar{x}_t + (I - A_3) x_t,$$
 (16)

$$s_1 = h + Hd + \varepsilon_1, \tag{17}$$

$$\bar{x}_1 = \bar{x}_{1,0} + Xd.$$
 (18)



$$x_t^* = c_{0,t} + C_{1,t}d + C_{2,t}s_t + C_{3,t}\bar{x}_t + \nu_t,$$
(19)

- c<sub>0,t</sub> is a K-dimensional vector,
- $C_{1,t}$  is a  $K \times L$  matrix,
- $C_{2,t}$  and  $C_{3,t}$  are K-dimensional diagonal matrices, and
- $\nu_t$  is a *K*-dimensional vector of rescaled preference shocks (i.e., I can write  $\nu_t = M_t \tilde{\nu}_t$ , where  $M_t$  is a *K*-dimensional diagonal matrix).



## **III. Estimation Strategy**



## A. Identification



- The scale parameters of skills are not identified because observed variables (i.e., wage and task complexity) are the product of unobserved skills and unknown parameters such as the returns to skills.
- Observed high wages can be rationalized by either a large amount of skills or high returns to skills.
- The location parameters of skills are also not identified because no natural measures of skills exist.
- Normalize skills by assuming that the unconditional mean and variance of initial skills are 0 and 1, respectively.



- Specifically,  $E(s_1) = h + HE(d) = 0$ ,
- diag[Var(s<sub>1</sub>)] = diag[HE(dd')H' + Σ<sub>ε1</sub>] = diag[I], where 0 is a vector of zeros,
- I is an identity matrix, and
- diag is an operator that converts a matrix into a vector that consists of diagonal elements of the matrix.
- Marginal effect of x on ln w (i.e., ∂ ln w/∂x = p<sub>1</sub> + P<sub>3</sub>s) can be consistently estimated regardless of how I normalize the location parameter.
- The covariances of the initial unobserved skills and skill shocks (i.e., off-diagonal elements of  $\Sigma_{\varepsilon}$  and  $\Sigma_{\varepsilon_1}$ ) are identified by the conditional wage variance given tasks.



- The parameters for job preferences are, in principle, identified by the occupational choices characterized by the policy function.
- However, they are not separately identified because the number of these parameters in the job preference equation (9) is greater than the number of the parameters in the policy function (19).
- To see this unidentifiability, consider the optimal choice of occupation at the terminal period *T*.
- The optimal task complexity  $x_T^*$  is given by

$$\begin{aligned} x_T^* &= -(1/2)(G_3 + G_4)^{-1}[g_0 + G_1d + (P_3 + G_2)s_T - 2G_4\bar{x}_T + \tilde{\nu}_T] \\ &\equiv c_{0,T} + C_{1,T}d + C_{2,T}s_T + C_{3,T}\bar{x}_T + \nu_T. \end{aligned}$$



#### **B. Kalman Filter**



- Use Kalman filter to calculate the likelihood.
- The Kalman filter is an algorithm used to recursively estimate the distribution of unobserved state variables (i.e., skills) from observed noisy signals (i.e., the task complexities of occupations and wages).
- Suppose that skills are normally distributed given task complexity x<sub>t</sub> and wages w<sub>t</sub> up to year t − 1, the initial work habit x
  <sub>1</sub>, and fixed worker characteristics d.



The conditional mean and variance of skills are

$$E(s_t|x_1, w_1, \dots, x_{t-1}, w_{t-1}; \bar{x}_1, d) \equiv E(s_t|Y_{t-1})$$
(20)  
$$\equiv \hat{s}_{t|t-1},$$
(21)

$$Var(s_t | x_1, w_1, \dots, x_{t-1}, w_{t-1}; \bar{x}_1, d) \equiv Var(s_t | Y_{t-1})$$
(22)  
$$\equiv \sum_{t|t-1}^{s},$$
(23)

• where  $T_{t-1}$  summarizes all the information up to year t-1.



• The conditional mean and variance of  $x_t$  given  $Y_t$  are

$$E(x_t|Y_{t-1}) = c_{0,t} + C_{1,t}d + C_{2,t}\hat{s}_{t|t-1} + C_{3,t}\bar{x}_t, \qquad (24)$$

$$Var(x_t|Y_{t-1}) = C_{2,t} \Sigma_{t|t-1}^s C_{2,t}' + \Sigma_{\nu}.$$
 (25)



- I then update the conditional distribution of skills using the task complexity in the current period  $x_t$  so that  $E(s_t|Y_{t-1}, x_t) = \hat{s}_{t|t-1} + \sum_{t|t-1}^{s} C'_{2,t} (C_{2,t} \sum_{t|t-1}^{s} C'_{2,t} + \sum_{\nu})^{-1} \hat{\nu}_t$ , (26)  $Var(s_t|Y_{t-1}, x_t) = \sum_{t|t-1}^{s} -\sum_{t|t-1}^{s} C'_{2,t} (C_{2,t} \sum_{t|t-1}^{s} C'_{2,t} + \sum_{\nu})^{-1} C_{2,t} \sum_{t|t-1}^{s} (27)$ 
  - where  $\hat{\nu}_t$  is a vector of residuals and  $\hat{\nu}_t = x_t E(x_t | Y_{t-1})$ .



- Notice that the log wage is a linear function of normal random variables given information up to t 1 and the current occupational tasks x<sub>t</sub>.
- Thus, the log wage is also normally distributed given and  $Y_{t-1}$  and  $x_t$ .
- The conditional mean and variance of the log wage are  $E(\ln w_t | Y_{t-1}, x_t) = p_0 + p'_1 x_t + [p_2 + P'_3 x_t]' E(s_t | Y_{t-1}, x_t),$  (28)  $Var(\ln w_t | Y_{t-1}, x_t) = [p_2 + P'_3 x_t]' Var(s_t | Y_{t-1}, x_t) [p_2 + P'_3 x_t] + \sigma_{\eta}^2.$



(29)

• Again, I then update the conditional distribution of skills using the information obtained in the current period:

$$E(s_t|Y_{t-1}, x_t, w_t) = E(s_t|Y_{t-1}, x_t) + \operatorname{Var}(s_t|Y_{t-1}, x_t)[p_2 + P'_3 x_t][\operatorname{Var}(\ln w_t|Y_{t-1} x_t)]^{-1}\hat{\eta}_t,$$
(30)



$$\begin{aligned} \mathsf{Var}(s_t | Y_{t-1}, x_t, w_t) \\ &= \mathsf{Var}(s_t | Y_{t-1}, x_t) \\ - \mathsf{Var}(s_t | Y_{t-1}, x_t) [p_2 + P_3' x_t] [\mathsf{Var}(\mathsf{In} \ w_t | Y_{t-1} x_t)]^{-1} [p_2 + P_3' x_t]' \mathsf{Var}(s_t | Y_{t-1}, x_t), \end{aligned}$$
(31)

• where  $\hat{\eta}_t$  is the log wage residual, and

• 
$$\hat{n}_t = \ln w_t - E(\ln w_t | Y_{t-1}, x_t).$$



- Finally, I calculate the conditional distribution of skills in year t + 1, given information up to year t, using the skill transition equation (see eq. (6)).
- Because skills in year t + 1 are linear in current skills and task complexity, they are also normally distributed with mean and variance:

$$\hat{s}_{t+1|t} = DE(s_t|Y_{t-1}, x_t, w_t) + a_0 + A_1x_t + A_2d,$$
 (32)

$$\Sigma_{t+1|t}^{s} = D \operatorname{Var}(s_{t}|Y_{t-1}, x_{t}, w_{t})D + \Sigma_{\varepsilon}.$$
(33)



- When the wage is missing in the data, it is integrated out to construct the likelihood.
- Given the linear skill transition equation (6), the conditional distribution of  $s_{t+1}$  given  $Y_{t-1}$  and  $x_t$  is normal with mean and variance

$$E(s_{t+1}|Y_{t-1}, x_t) = Ds_t + a_0 + A_1 x_t + A_2 d$$
  
=  $D\hat{s}_{t|t-1} + a_0 + A_1 x_t + A_2 d$ , (34)  
 $Var(s_{t+1}|Y_{t-1}, x_t) = D\Sigma_{t|t-1}^s D + \Sigma_{\varepsilon}$ , (35)

• which replace equations (32) and (33) in this case.



• The likelihood contribution of individual *i* is

$$I(w_{i1}, x_{i1}, \dots, w_{iT_i}, x_{iT_i} | \bar{x}_{i1}, d_i) = I(x_{i1} | \bar{x}_{i1}, d_i) I(w_{i1} | x_{i1}; \bar{x}_{i1}, d_i) \times \dots \times I(x_{iT_i} | Y_{iT_i-1}) I(w_{i1} | Y_{iT_i-1}, x_{iT_i}).$$
(36)

• The likelihood for the whole sample consisting of *N* individuals is given by

$$I = \prod_{i=1}^{N} I(w_{i1}, x_{i1}, \dots, w_{iT_i}, x_{iT_i} | \bar{x}_{i1}, d_i).$$
(37)



# IV. Data



## A. Dictionary of Occupational Titles



	Cognitive Task		Motor Task			
	Mean	SD	Mean	SD	No. Observations	
Professional	.85	.14	.45	.33	7,522	
Manager	.79	.15	.21	.21	5,538	
Sales	.57	.17	.23	.15	3,748	
Clerical	.49	.16	.56	.22	9,270	
Craftsmen	.52	.20	.82	.20	6,557	
Operatives	.20	.18	.58	.20	5,824	
Transport	.28	.15	.63	.10	1,774	
Laborer	.15	.16	.46	.13	2,818	
Farmer	.68	.19	.78	.14	1,117	
Farm laborer	.18	.19	.53	.16	882	
Service	.32	.22	.44	.24	6,834	
Household service	.20	.11	.24	.23	1,469	
All occupations	.49	.29	.50	.29	53,353	

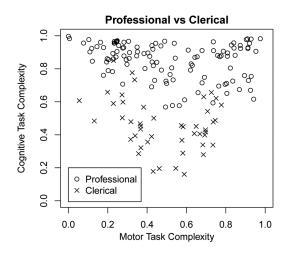
#### Table 1: Task Complexity by Occupation at One-Digit Classification

*Note*: Sample consists of all working individuals in the 1971 April Current Population Survey augmented with occupational characteristics variables from the revised fourth edition of the Dictionary of Occupational Titles (1991). Sample size is 53,353. Task complexity measures are percentile scores divided by 100.

# **B.** National Longitudinal Survey of Youth 1979



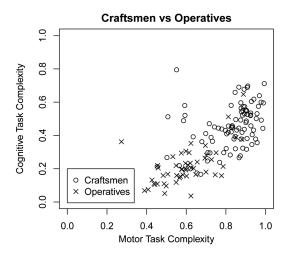
#### Figure 1: Task complexity comparison



Task complexity measures are percentile scores divided by 100.

Source: 1971 April Current Population Survey augmented with occupational characteristics variables from the revised fourth edition of the Dictionary of Occupational Titles (1991).

#### Figure 1: Task complexity comparison



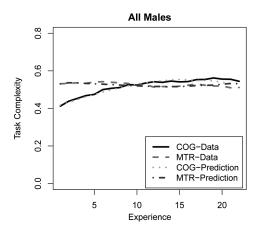
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Source: 1971 April Current Population Survey augmented with occupational characteristics (TY OF variables from the revised fourth edition of the Dictionary of Occupational Titles (1991): A C O

## **C.** Career Progression Patterns

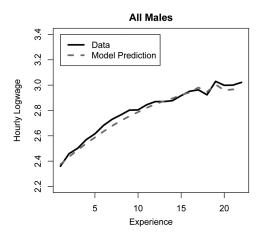


Figure 2: Model fit (all men)



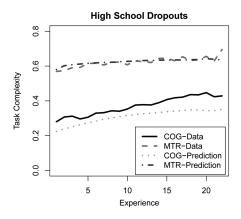
COG = profiles of cognitive task complexity; MTR = profiles of motor task complexity.Source: Author's estimates from the National Longitudinal Survey of Youth 1979–2000.Sample consists of 325 high school dropouts, 1,009 high school graduates, and 1,083 college workers. Color version available as an online enhancement.

## Figure 2: Model fit (all men)



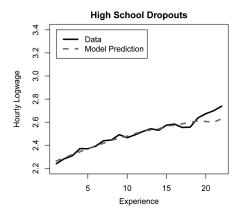
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#### Figure 3: Model fit (high school dropouts)



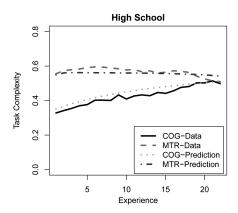
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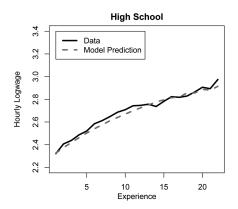
COG = profiles of cognitive task complexity; MTR = profiles of motor task complexity. *Source*: Author's estimates from the National Longitudinal Survey of Youth 1979–2000. Sample includes 325 high school dropouts. Color version available as an online enhancement.

### Figure 4: Model fit (high school workers)



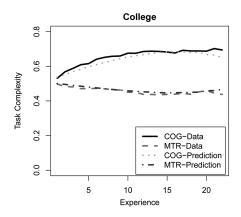
COG = profiles of cognitive task complexity; MTR = profiles of motor task complexity. *Source*: Author's estimates from the National Longitudinal Survey of Youth 1979–2000. Sample includes 1,009 high school graduates. Color version available as an online enhancement.

### Figure 4: Model fit (high school workers)



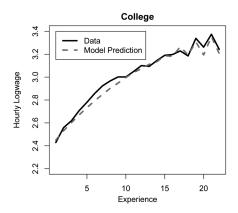
COG = profiles of cognitive task complexity; MTR = profiles of motor task complexity. Source: Author's estimates from the National Longitudinal Survey of Youth 1979–2000. Sample includes 1,009 high school graduates. Color version available as an online enhancement.

#### Figure 5: Model fit (college workers)



COG = profiles of cognitive task complexity; MTR = profiles of motor task complexity. *Source*: Author's estimates from the National Longitudinal Survey of Youth 1979–2000. Sample includes 1,083 college workers. Color version available as an online enhancement.

#### Figure 5: Model fit (college workers)



COG = profiles of cognitive task complexity; MTR = profiles of motor task complexity. Source: Author's estimates from the National Longitudinal Survey of Youth 1979–2000. Sample includes 1,083 college workers. Color version available as an online enhancement.



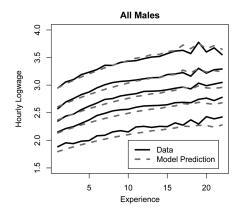
## V. Estimation Results



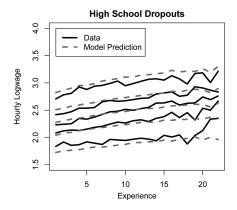
# A. Model Fit



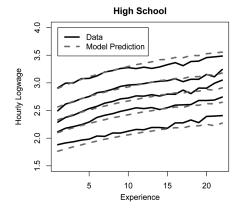
Figure 6: Model fit: wage growth by percentile (all men and high school dropouts)



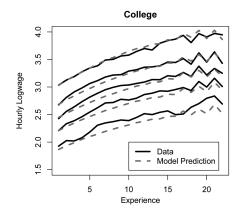
Tenth, 30th, 50th, 70th, and 90th percentiles of hourly log wages are plotted over time. *Source*: Author's estimates from the National Longitudinal Survey of Youth 1979–2000. Sample consists of 325 high school dropouts, 1,009 high school graduates, and 1,083 college, workers. Color version available as an online enhancement. Figure 6: Model fit: wage growth by percentile (all men and high school dropouts)



Tenth, 30th, 50th, 70th, and 90th percentiles of hourly log wages are plotted over time. *Source*: Author's estimates from the National Longitudinal Survey of Youth 1979–2000. Sample consists of 325 high school dropouts, 1,009 high school graduates, and 1,083 college workers. Color version available as an online enhancement. Figure 7: Model fit: wage growth by percentile (high school and college workers)



Tenth, 30th, 50th, 70th, and 90th percentiles of hourly log wages are plotted over time. *Source*: Author's estimates from the National Longitudinal Survey of Youth 1979–2000. Sample includes 1,009 high school graduates and 1,083 college workers. Color version Figure 7: Model fit: wage growth by percentile (high school and college workers)



Tenth, 30th, 50th, 70th, and 90th percentiles of hourly log wages are plotted over time. *Source*: Author's estimates from the National Longitudinal Survey of Youth 1979–2000. Sample includes 1,009 high school graduates and 1,083 college workers. Color version international control of the second seco

## Table 2: Wage Equation

Notation	Estimate	SE
$p_{\circ}$	2.294	.014
$p_{1}(1)$	.015	.012
$p_1(2)$	.106	.019
$p_2(1)$	.646	.065
$p_2(2)$	.575	.078
$P_{3}(1, 1)$	.108	.010
$P_{3}(2, 2)$	.047	.012
$P_{3}(2, 2) \sigma_{\eta}^{2}$	.061	.001

Source: National Longitudinal Survey of Youth 1979–2000. Sample consists of 2,417 men. Note: Parameter estimates are for the wage equation  $\ln w_t = p_0 + p'_1 x_t + (p_2 + P'_3 x_t)'s_t + \eta_t$ , where  $\eta_t \sim N(0, \sigma_{\eta}^2)$ . The first element of vectors and (1, 1) element of matrices are for cognitive skills, and the second element of vectors and (2, 2) element of matrices are motor skills.

## **B.** Parameter Estimates



# 1. Wage Equation



# 2. Skill Transition



### Table 3: Transition Equation and Initial Conditions for Skills

Notation	Estimate	SE	
D(1, 1)	.925	.002	
D(2, 2)	.911	.008	
$a_0(1)$	309	.034	
$a_0(2)$	.227	.048	
$A_1(1, 1)$	.057	.009	
$A_1(2, 2)$	.101	.036	
$A_2(1, 1), AFQT$	.071	.032	
A,(1, 2), Edu	.026	.003	
A,(1, 3), Black	.015	.025	
$A_2(1, 4)$ , Hispanic	040	.026	
$A_2(2, 1)$ , AFQT	044	.044	
A2(2, 2), Edu	020	.005	
A <sub>2</sub> (2, 3), Black	036	.033	
A <sub>2</sub> (2, 4), Hispanic	.050	.035	
$\Sigma_{e}(1,1)$	.156	.019	
$\Sigma_{e}(2, 1)$	125	.019	
$\Sigma_{r}(2,2)$	.145	.021	
H(1, 1), AFQT	.514	.410	
H(1, 2), Edu	.184	.040	
H(1, 3), Black	.500	.327	
H(1, 4), Hispanic	237	.344	
H(2, 1), AFQT	051	.469	
H(2, 2), Edu	202	.046	
H(2, 3), Black	638	.366	
H(2, 4), Hispanic	.371	.388	
$\Sigma_{e}(1,1)$	.654	.075	
$\Sigma_{\epsilon_1}^{\epsilon_1}(2, 1)$	519	.080	
$\Sigma_{\epsilon_1}^{\epsilon_1}(2,2)$	.685	.089	



## 3. Job Preference



## **VI.** Discussion



# A. Do Cognitive and Motor Skills Account for Wage Variance?



Table 4: Job	Preference	and	Work	Habits
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Notation	Estimate	SE
$g_0(1)$	.075	.119
$g_0(2)$	171	.278
$g_{0}(2)$ $G_{1}(1, 1), AFQT$	.295	.074
$G_1(1, 2), Edu$	.038	.009
$G_1(1, 3)$ , Black	190	.056
$G_1(1, 4)$ , Hispanic	.055	.063
$G_1(2, 1), AFQT$	.047	.172
$G_1(2, 2), Edu$	.043	.020
$G_1(2, 3)$ , Black	.142	.134
$G_1(2, 4)$ , Hispanic	208	.141
$G_2(1, 1)$	.037	.020
$G_{2}(2, 2)$	.326	.030
$G_{4}(1, 1)$	-13.048	.780
$G_4(2, 2)$	146	.045
$A_{3}(1, 1)$	.475	.007
$A_{3}(2, 2)$	.000	.074
$\bar{x}_{1,0}(1)$	218	.024
X(1, 1), AFQT	.152	.019
X(1, 2), Edu	.040	.002
X(1, 3), Black	017	.015
X(1, 4), Hispanic	.050	.017
$\bar{x}_{1,0}(2)$	.255	.198
X(2, 1), AFQT	.158	.155
X(2, 2), Edu	.007	.017
X(2, 3), Black	.219	.125
X(2, 4), Hispanic	.259	.131 📑 THE UNIVERS
		CHICA

The wage can be written as

$$\ln w_t = p_0 + [p_1^C x_t^C + p(x_t)^C s_t^C] + [p_1^M x_t^M + p(x_t)^M s_t^M] + \eta_t,$$

- where superscript C is for cognitive skills, and
- M is for motor skills.
- I call the term  $p_1^C x_t^C + p(x_t)^C s_t^C$  the cognitive skill component and the term  $p_1^M x_t^M + p(x_t)^M s_t^M$  the motor skill component.



	Cognitive	Motor	Covariance	Error	Total
Year 1:					
All men	.482	.378	715	.061	.205
High school dropouts	.353	.339	565	.061	.188
High school	.321	.287	471	.061	.198
College	.421	.323	600	.061	.204
Year 10:					
All men	.913	.546	-1.226	.061	.293
High school dropouts	.565	.470	871	.061	.224
High school	.502	.399	729	.061	.233
College	.755	.458	997	.061	.277
Year 20:					
All men	1.167	.617	-1.484	.061	.360
High school dropouts	.652	.520	991	.061	.242
High school	.571	.431	813	.061	.249
College	.934	.506	-1.173	.061	.328

#### Table 5: Log Wage Variance Decomposition

*Note*: Author's estimates from the National Longitudinal Survey of Youth 1979–2000. Sample consists of 325 high school dropouts, 1,009 high school graduates, and 1,083 college workers.

		Homogeneous			
Year	Benchmark	Preference	Initial Skills	Learning Ability	All
1	.206	.204	.061	.206	.061
10	.292	.260	.234	.241	.190
20	.359	.297	.335	.257	.232

#### Table 6: Log Wage Variance When Initial Conditions Are Homogeneous

*Note*: Author's estimates from the National Longitudinal Survey of Youth 1979–2000. Sample consists of 2,417 men.



# Table 7: College/High School Log Wage Gaps When Initial Conditions Are Homogeneous

Year	Benchmark	Preference	Initial Skills	Leaning Ability
1	.126	.107	.011	.126
10	.340	.231	.283	.155
20	.460	.308	.432	.167

*Note*: Author's estimates from the National Longitudinal Survey of Youth 1979–2000. Sample consists of 2,417 men.



Year	All Men	High School Dropouts	High School	College
Cognitive skills:				
1	.000	813	269	.498
10	.631	650	.206	1.405
20	.996	539	.489	1.923
Motor skills:				
1	.000	.731	.240	448
10	066	.871	.240	637
20	108	.950	.238	750

### Table 8: Mean Skill Profiles by Education

*Note*: Author's estimates from the National Longitudinal Survey of Youth 1979–2000. Sample consists of 325 high school dropouts, 1,009 high school graduates, and 1,083 college workers.



#### B. The Growth of Skills and Wages



		Benchmark				All Men	
Years since Entry	Dropouts (1)	High School (2)	College (3)	All Men (4)	CF 1 (5)	CF 2 (6)	
Cognitive skills:							
5	.068	.209	.416	.282	.276	.271	
10	.124	.362	.716	.487	.472	.454	
15	.166	.472	.927	.634	.612	.580	
20	.197	.549	1.074	.737	.710	.665	
Motor skills:							
5	.061	.003	069	021	026	027	
10	.098	.003	120	038	045	044	
15	.124	.002	156	052	059	055	
20	.141	.000	183	063	071	062	
Total:							
5	.129	.212	.347	.261	.249	.244	
10	.222	.365	.596	.448	.427	.410	
15	.291	.474	.771	.582	.552	.525	
20	.337	.549	.892	.674	.639	.603	

#### Table 9: Accumulated Wage Growth by Skill Type and Education

*Note*: Author's estimates from the National Longitudinal Survey of Youth 1979–2000. Sample consists of 325 high school dropouts, 1,009 high school graduates, and 1,083 college workers. CF = counterfactual.

### **VII.** Conclusion



### Appendix A Identification



- Denote the *k*th element of a vector z by z(k).
- Similarly, denote the (k, l) element of a matrix Z by Z(k, l).
- The wage equation can be rewritten as

$$\ln w = p_0 + \sum_{k=1}^{K} p_1(k) x(k) + \sum_{k=1}^{K} [p_2(k) + P_3(k,k) x(k)] s(k) + \eta.$$

• To simplify the discussion, assume that skills are given by

$$s(k) = \alpha_1(k)z_1(k) + \alpha_2(k)z_2 + \tilde{s}(k),$$

- where z<sub>1</sub> is a vector of variables of which the kth element affects the kth skills only (i.e., task complexity of the past job),
- *z*<sub>2</sub> is a scalar variable that can affect all types of skills (i.e., education, AFQT score, race), and
- $\tilde{s}$  is a vector of unobserved components of skills.

Substituting the skill equation, I rewrite the wage equation as

$$\ln w = p_0 + \sum_{k=1}^{K} p_1(k) x(k) + \sum_{k=1}^{K} [p_2(k) + P_3(k, k) x(k)] [z_1(k) + \alpha_2(k) z_2 + \tilde{s}(k)] + \eta$$

Rearranging terms, I have

$$\ln w = p_0 + \sum_{k=1}^{K} [p_1(k) + P_3(k, k)\tilde{s}(k)]x(k) + \sum_{k=1}^{K} p_2(k)z_1(k) + \sum_{k=1}^{K} p_2(k)\alpha_2(k)z_2 + \sum_{k=1}^{K} P_3(k, k)x(k)z_1(k) + \sum_{k=1}^{K} P_3(k, k)\alpha_2(k)x(k)z_2 + \left(\sum_{k=1}^{K} p_2(k)\tilde{s}(k) + \eta\right)$$



• Consider the conditional mean of the log wage:

$$E(\ln w|x, z_1, z_2) = p_0 + \sum_{k=1}^{K} p_1(k)x(k) + \sum_{k=1}^{K} p_2(k)z_1(k) + \sum_{k=1}^{K} p_2(k)\alpha_2(k)z_2 + \sum_{k=1}^{K} P_3(k, k)x(k)z_1(k) + \sum_{k=1}^{K} P_3(k, k)\alpha_2(k)x(k)z_2.$$

 The variation of x, z<sub>1</sub>, z<sub>2</sub>, xz<sub>1</sub>, xz<sub>2</sub>, and the conditional mean of the log wage identifies the parameters p<sub>0</sub>, p<sub>1</sub>, p<sub>2</sub>, P<sub>3</sub>, and α<sub>2</sub>.



• The variance and covariance of unobserved skills are identified by the conditional variance of log wage, which is given by

$$V(\ln w | x, z_1, z_2) = \sum_{k=1}^{K} \sum_{l=1}^{K} [p_2(k) + P_3(k, k) x(k)] \\ \times [p_2(l) + P_3(l, l) x(l)] \operatorname{Cov}[\tilde{s}(k), \tilde{s}(l)] + \sigma_{\eta}^2.$$

Given that p<sub>2</sub> and P<sub>3</sub> are identified by the conditional mean log wage, the variance and covariance of unobserved skills are identified by the variation of x(k)x(l) and the conditional variance.



#### Appendix B Robustness: Normally Distributed Task Indexes



Table 10: Estimates of Correlation Coefficients for Unobserved Variables	Table 10	: Estimates of	Correlation	Coefficients fo	r Unobserved	Variables
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	Percentil	e Score	Normally Distributed Index		
Notation	Estimate	SE	Estimate	SE	
Corr $(\varepsilon_1(1), \varepsilon_1(2))$	7757	.0498	5751	.0779	
Corr $(\varepsilon(1), \varepsilon(2))$	8332	.0351	6865	.0587	
Corr $(\nu(1), \nu(2))$	0295	.0038	.0102	.0038	

Source: National Longitudinal Survey of Youth 1979–2000. Sample consists of 2,417 men. Note: Variables  $\varepsilon_1$ ,  $\varepsilon$ , and  $\nu$  are initial unobserved skills, skill shocks, and preference shocks, respectively. For each variable, the first element is for cognitive skills/tasks, and the second element is for motor skills/tasks.



## Table 11: Log Wage Variance Decomposition for Normally Distributed Task Indexes

	Cognitive	Motor	Covariance	Error	Total
Year 1:					
All men	.466	.301	622	.061	.206
Dropouts	.287	.219	374	.061	.193
High school	.221	.136	219	.061	.198
College	.347	.209	412	.061	.205
Year 10:					
All men	.818	.393	978	.061	.293
Dropouts	.434	.281	547	.061	.228
High school	.329	.186	341	.061	.235
College	.589	.276	654	.061	.272
Year 20:					
All men	1.025	.435	-1.165	.061	.355
Dropouts	.499	.304	617	.061	.247
High school	.370	.202	382	.061	.251
College	.720	.304	771	.061	.313

*Note*: Author's estimates from the National Longitudinal Survey of Youth 1979–2000. Sample consists of 325 high school dropouts, 1,009 high school graduates, and 1,083 college workers.



# Table 12: Accumulated Wage Growth by Skill Type and Education for Normally Distributed Task Indexes

Years since Entry All Me		High School Dropouts	High School	College	
Cognitive skills:					
5	.272	.086	.210	.387	
10	.466	.153	.361	.662	
20	.699	.237	.542	.983	
Motor skills:					
5	008	.041	.006	036	
10	015	.066	.009	062	
20	030	.093	.008	100	
Total:					
5	.264	.128	.216	.352	
10	.451	.219	.370	.599	
20	.669	.328	.550	.884	

*Note*: Author's estimates from the National Longitudinal Survey of Youth 1979–2000. Sample consists of 325 high school dropouts, 1,009 high school graduates, and 1,083 college workers.

