

# Sampling Plans and Initial Condition Problems For Continuous Time Duration Models

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## Sampling plans and initial condition problems: Duration Models

Consider a random sample of unemployment spells in progress. For sampled spells, one of the following duration times may be observed:

- time in state up to sampling date ( $T_b$ ) (recall of time spent)
- time in state after sampling date ( $T_a$ ) (sampling forward)
- total time in completed spell observed at origin of sample ( $T_c = T_a + T_b$ )

Duration of spells beginning after the origin date of the sample, denoted  $T_d$ , are not subject to initial condition problems. The intake rate at time  $-t_b$  (assuming sample occurs at time 0: the proportion of the population entering a spell at  $-t_b$ .

### Assume:

- A time homogenous environment, i.e. constant intake rate,  $k(-t_b) = k, \forall b$
- A model without observed or unobserved explanatory variables.
- No right censoring, so  $T_c = T_a + T_b$
- Underlying distribution  $f(x)$  is nondefective
- $m = \int_0^\infty (x) dx < \infty$

The proportion of the population experiencing a spell at  $t = 0$ , the origin date of the sample, is

$$\begin{aligned} P_0 &= \int_0^{\infty} k(-t_b)(1 - F(t_b))dt_b = k \int_0^{\infty} (1 - F(t_b))dt_b \\ &= k \left[ t_b(1 - F(t_b)) \Big|_0^{\infty} - \int_0^{\infty} t_b d(1 - F(t_b)) \right] \\ &= k \int_0^{\infty} t_b f(t_b) dt_b = km \end{aligned}$$

where  $1 - F(t_b)$  is the probability the spell lasts from  $-t_b$  to 0 (or equivalently, from 0 to  $-t_b$ ).

So the density of a spell of length  $t_b$  interrupted at the beginning of the sample ( $t = 0$ ) is

$$\begin{aligned} g(t_b) &= \frac{\text{proportion surviving til } t = 0 \text{ from batch } t_b}{\text{total surviving til } t = 0} \\ &= \frac{k(-t_b)(1 - F(t_b))}{P_0} = \frac{1 - F(t_b)}{m} \neq f(t_b) \end{aligned}$$

*Notice:*  $g$  is the distribution of  $T_b$  in the population constructed by sampling rule of source population.

Distinguish from  $F$  : *cdf* of the true population.  $G$  : *cdf* of the sampled spells.

The probability that a spell lasts until  $t_c$  given that it has lasted from  $-t_b$  to 0, is the conditional density of  $t_c$  given  $0 < t_b < t_c$ .

$$f(t_c | t > t_b > 0) = \frac{f(t_c)}{1 - F(t_b)}; t_c \geq t_b \geq 0$$

So the density of a spell **in the sampled population** that lasts,  $t_c$  is

$$\begin{aligned} g(t_c) &= \int_0^{t_c} f(t_c | t \geq t_b) f(t \geq t_b) dt_b \\ &= \int_0^{t_c} \frac{f(t_c)}{m} dt_b = \frac{f(t_c)t_c}{m} \end{aligned}$$

Likewise, the density of a sampled spell that lasts until  $t_a$  is

$$\begin{aligned}g(t_a) &= \int_0^\infty f(t_a + t_b | t_b) \Pr(t \geq t_a \geq 0) dt_b \\ &= \int_0^\infty \frac{f(t_a + t_b)}{m} dt_b \\ &= \frac{1}{m} \int_{t_a}^\infty f(t_b) dt_b \\ &= \frac{1 - F(t_a)}{m}\end{aligned}$$

(Stationarity, mirror images have same densities). So the functional form of  $f(t_b) = f(t_a)$ : Consequences of stationarity.



Some useful results that follow from this model:

- If  $f(t) = \theta e^{-t\theta}$ , then  $g(t_b) = \theta e^{-t_b\theta}$  and  $g(t_a) = \theta e^{-t_a\theta}$ .

**Proof:**

$$f(t) = \theta e^{-t\theta} \rightarrow m = \frac{1}{\theta},$$

$$F(t) = 1 - e^{-t\theta} \rightarrow g(t_a) = \frac{1 - F(t)}{m} = \theta e^{-t\theta}$$

- $E(T_a) = \frac{m}{2} \left(1 + \frac{\sigma^2}{m^2}\right).$

**Proof:**

$$\begin{aligned} E(T_a) &= \int t_a f(t_a) dt_a = \int t_a \frac{1 - G(t_a)}{m} dt_a \\ &= \frac{1}{m} \left[ \frac{1}{2} t_a^2 (1 - F(t_a)) \Big|_0^\infty - \int \frac{1}{2} t_a^2 d(1 - F(t_a)) \right] \\ &= \frac{1}{m} \int \frac{1}{2} t_a^2 F(t_a) dt_a = \frac{1}{2m} [\text{var}(t_a) + E^2(t_a)] \\ &= \frac{1}{2m} [\sigma^2 + m^2] \end{aligned}$$

- $E(T_b) = \frac{m}{2} \left(1 + \frac{\sigma^2}{m^2}\right)$ . **Proof:** See proof of Proposition 2.
- $E(T_c) = m \left(1 + \frac{\sigma^2}{m^2}\right)$ . **Proof:**

$$E(T_c) = \int \frac{t_c^2 F(t_c)}{m} dt_c = \frac{1}{m} (\text{var}(t_c) + E^2(t_c))$$

$$\rightarrow E(T_c) = 2E(T_a) = 2E(T_b), E(T_c) > m \text{ unless } \sigma^2 = 0$$

## Some Additional Results:

$$h(t) = \text{hazard} : h(t) = \frac{f(t)}{1 - F(t)}.$$

- $h'(t) > 0 \rightarrow E(T_a) = E(T_b) < m$ . **Proof:** See Barlow and Proschan.
- $h'(t) < 0 \rightarrow E(T_a) = E(T_b) > m$ . **Proof:** See Barlow and Proschan.

## Examples

## Specification of the Distribution

### Weibull Distribution

- Parameters:  $\lambda > 0, k > 0$
- Probability Density Function (PDF):

$$\frac{\lambda}{k} \left(\frac{t}{\lambda}\right)^{k-1} \exp\left(-\left(\frac{t}{\lambda}\right)^k\right)$$

- Cumulative Density Function:

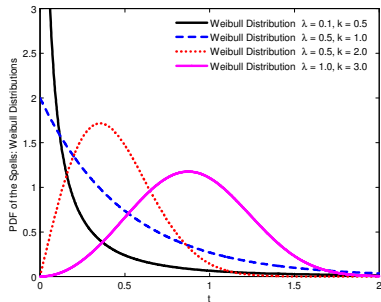
$$1 - \exp\left(-\left(\frac{t}{\lambda}\right)^k\right)$$

- Set of Parameters:

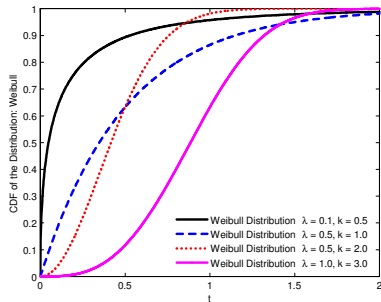
$$\left( \begin{array}{l} \lambda_1, k_1 = 0.5 \\ \lambda_2, k_1 = 1.0 \\ \lambda_3, k_1 = 2.0 \\ \lambda_3, k_1 = 3.0 \end{array} \right), \text{ respectively}$$

# Basic Distribution Graphs

PDF for Weibull Distribution

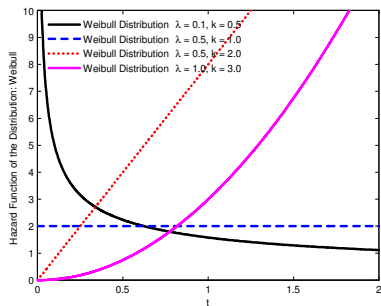


CDF of Weibull Distribution

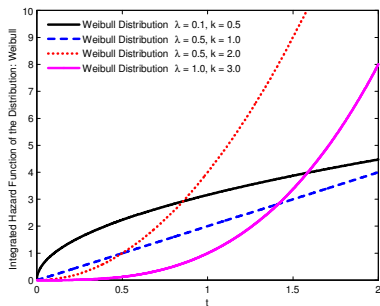


# Basic Duration Graphs

## Hazard Function for Weibull Distribution

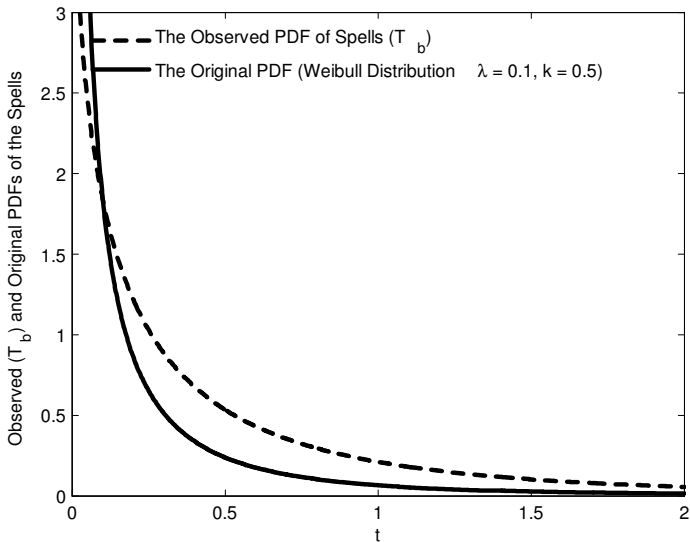


## Integrated Hazard Function for Weibull

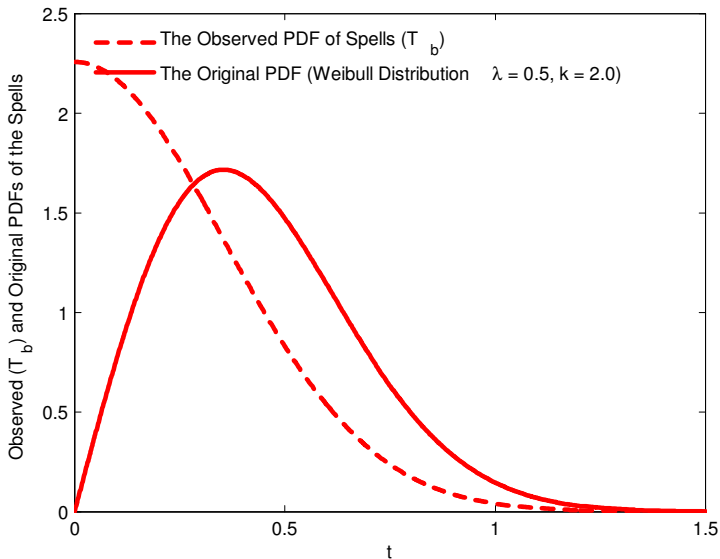




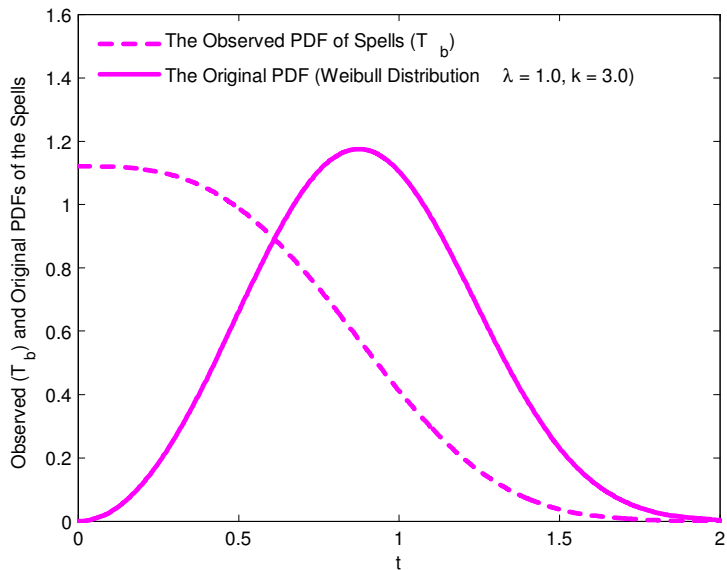
## Observed and Original Distribution for $T_b$ (Example 1)



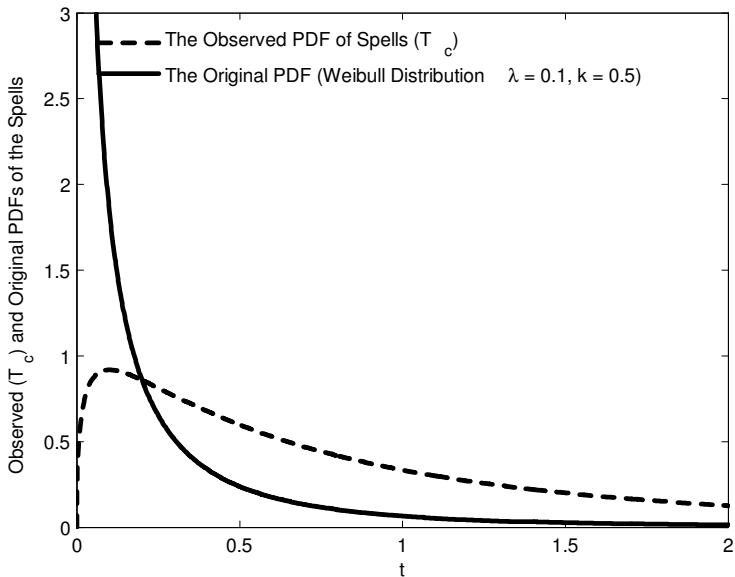
## Observed and Original Distribution for $T_b$ (Example 3)



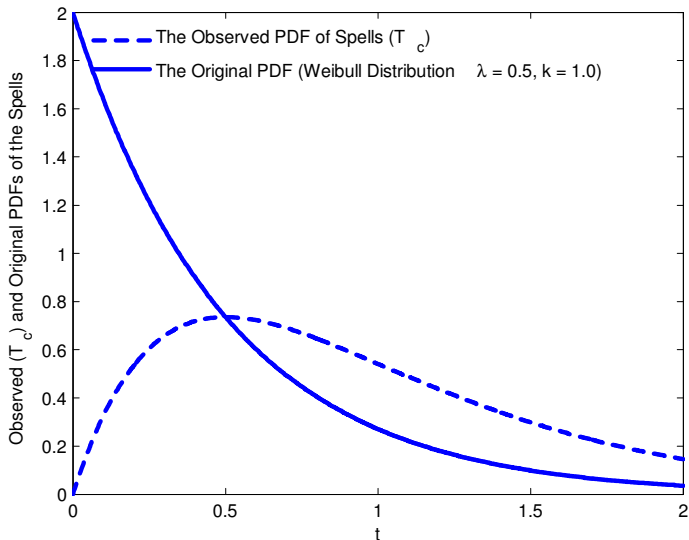
## Observed and Original Distribution for $T_b$ (Example 4)



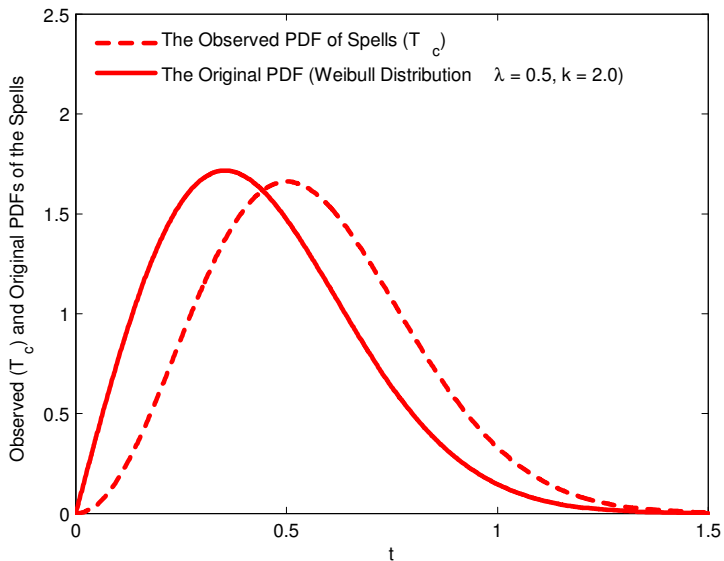
## Observed and Original Distribution for $T_c$ (Example 1)



## Observed and Original Distribution for $T_c$ (Example 2)



## Observed and Original Distribution for $T_c$ (Example 3)



## Observed and Original Distribution for $T_c$ (Example 4)

